STAT 8025

Lecture 13: Point Process

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Introduction

- Point processes: Locations are random
- ► Example: Tree locations
 - Ambrosia dumosa is a drought deciduous shrub with 20-60cm in height which is abundant on well drained soils below one thousand meter elevation. The data were collected within a hectare (100 × 100m²) area in the Colorado Desert in 1984. The data contains
 - ne data contains
 - locations 4358 Ambrosia dumosa trees
 - the height of the plant canopy
 - the length of the major axis of the plant canopy
 - the length of the minor axis of the plant canopy
 - the volume of the plant canopy.
- Examples: occurrence of earthquake, crime, forest fire
- \blacktriangleright We observe s_1, \ldots, s_n , and model n and s_1, \ldots, s_n as random.



Types of Point Patterns

- ► Completely random: $s_i | n \stackrel{iid}{\sim} Uniform(S)$
- Clustered: s_i's form clusters
- ▶ Inhomogeneous: Events more likely in some regions than others (sometimes it can be hard to distinguish this from clustered pattern)
- Regular: s_i's repulse each other



Ripley's K Function

- Ripley's K function is analogous to the variogram in geostatistics
- ▶ Let $d_{ij} = |s_i s_j|$ be the distance between s_i and s_j
- ► Ripley's empirical K function is

$$\hat{K}(t) = |\mathcal{S}| \frac{1}{n^2} \sum_{i \neq j} I(d_{ij} \leq t)$$

- estimating proportion that a pair of points are within t of each other
- but this is the empirical function; as variogram, different modes have different true K-functions and plotting \hat{K} will help choose a true model



Point Processes

▶ Let $S \subset \mathbb{R}^d$ be the domain. Let N(A) be the number of points in A for any $A \subset S$. Then, the distribution of N is given by

$$P_k[N(A_1)=n_1,\cdots,N(A_k)=n_k]$$

for any $A_1, \ldots, A_k \subset \mathcal{S}$ and $k \in \mathcal{N}^+$.

▶ The k-th order intensity function of N (if exists) is

$$\lambda_k(\mathsf{s}_1,\ldots,\mathsf{s}_k) = \lim_{|d\mathsf{s}_i| \to 0, i=1,\ldots,k} \left\{ \frac{E[N(d\mathsf{s}_1) \cdots N(d\mathsf{s}_k)]}{|d\mathsf{s}_1| \cdots |d\mathsf{s}_k|} \right\}$$

where s_i are distinct points in S.

▶ Both P_k and λ_k can be used for the distribution of N, but people focus on λ_k more.



► The pair correlation function:

$$g(\mathsf{s}_1,\mathsf{s}_2) = \frac{\lambda_2(\mathsf{s}_1,\mathsf{s}_2)}{\lambda(\mathsf{s}_1)\lambda(\mathsf{s}_2)}$$

- ► The mean function $\mu(A) = \int_A \lambda(s) ds$
- ► The covariance structure *N* is

$$Cov[N(A_1), N(A_2)] = \int_{A_1} \int_{A_2} [\lambda_2(s_1, s_2) - \lambda(s_1)\lambda(s_2)] ds_2 ds_1 + \int_{A_1 \cup A_2} \lambda(s) ds$$

K-function

- ▶ Suppose *N* is stationary. Let λ be the first-order intensity function.
- ► The K-function is defined as

$$K(t) = \frac{1}{\lambda} E$$
[number of extra events within distance of t

of a randomly chosen event]

► The L-function is

$$L(t) = \sqrt{\frac{K(t)}{\pi}}$$

ightharpoonup K(t) is more often used.



Stationarity

A spatial point process N is said to be strongly stationary if for any A_1, \ldots, A_k , the joint distribution of

$$N(A_1 + s), \ldots, N(A_k + s)$$

does not depend on s, where

$$A_i + s = \{s' + s : s' \in A_i\}$$

▶ If $\lambda(s)$ is constant and

$$\lambda_2(\mathsf{s}_1,\mathsf{s}_2) = \lambda_2(\mathsf{s}_1-\mathsf{s}_2)$$

N is called second-order stationary. In addition, if

$$\lambda_2(s_1, s_2) = \lambda_2(|s_1 - s_2|)$$

N is called isotropic.



Poisson Point Process

A poisson point process is derived if $N(A_1), \ldots, N(A_k)$ are independent Poisson random variable with mean $\mu(|A_1|), \ldots, \mu(|A_k|)$ if A_1, \ldots, A_k are disjoint subsets.

- ► The simplest case with $\lambda(s) = \lambda$ and this implies: $n \sim Poisson(\lambda|S|)$ and $s_1, \ldots, s_n|n \sim Uniform(S)$
 - ► Called the homogeneous Poisson process, $n \sim Poisson(\lambda|S|)$ and is stationary and sampled completely at random.
 - The sufficient statistics for the single parameter is n and $\hat{\lambda} = \frac{n}{|S|}$ is the MLE



- Inhomogeneous Poisson point process: $\lambda(s)$ spatially-varying
 - $\mu(A) = \int_A \lambda(s) ds$ so the total number of event $n \sim Poisson(\mu(S))$ Siven $n, s_1, \dots, s_n \stackrel{iid}{\sim} f(s)$ with

$$f(s) = \frac{\lambda(s)}{\int_{S} \lambda(t) dt}$$

Simulate an inhomogeneous Poisson point process?

- ► Let f be a pdf on s.
- ▶ We generate $n \sim Poisson(B)$
- ▶ Generate *n* observations idependently from f, s_1, \ldots, s_n
- ► Then $s_1, ..., s_n$ is a sample from Point point process with intensity

$$\lambda(\mathsf{s}) = Bf(\mathsf{s})$$



Thinned Poisson Point Process

- We have $\lambda(s) = Z(s)\lambda$ where $Z(s) \in [0,1]$
- Simulating a thinned Poisson process
 - 1. Generate $n \sim Poisson(\lambda |\mathcal{S}|)$
 - 2. Generate $s_1, \ldots, s_n \sim \textit{Uniform}(\mathcal{S})$
 - 3. Keep s_i wit probability $Z(s_i)$



Modeling Inhomogeneous Poisson Process

▶ We can model the log intensity function:

$$log(\lambda(s)) = \beta_0 + \sum_{i=1}^{p} X_i(s)\beta_i$$

► The likelihood is:

$$I(\beta) = \prod_{i=1}^{n} \frac{\lambda(s_i)}{\int_{\mathcal{S}} \lambda(s) ds} = \frac{\exp[\sum_{i=1}^{n} X(s_i)' \beta]}{[\int_{\mathcal{S}} \lambda(s) ds]^n}$$

- ► The integration is difficult to compute.
- In practice, this integration is often approximated with a sum: Partition S into B_1, \ldots, B_m and the integration is approximated with $\sum_{i=1}^m \exp(X_i'\beta)$
- ► Reading assignment: Hessellund, Xu, Guan & Waagepetersen (2021) Semiparametric Multinomial Logistic Regression for Multivariate Point Pattern Data, Journal of the American Statistical

Association, DOI: 10.1080/01621459.2020.1863812.

Clustering

- ► An intuitive clustering model is to build it based on a parent Poisson point process:
 - ▶ The parent process is a Poisson point process, and we get v_1, \ldots, v_m follow a homogeneous Poisson point process. These m locations are centers of clusters.
 - For cluster j, j = 1, ..., m, we assume:

$$n_j \sim Poisson(\lambda_0)$$
 $s_{ji} \sim N(v_j, \sigma^2 I), ext{ for } i = 1, \dots, n_j$

▶ It can be fit using EM algorithm or MCMC.



Cox Process

Log-Gaussian Cox Process:

$$log(\lambda(s)) = X(s)'\beta + \theta(s)$$

 $\theta(s) \sim GP(\mu(\cdot), C(\cdot, \cdot))$

- Likelihood is difficult to deal with due to the integration... approximation is needed.
- ► Reading assignment:
 - Chakraborty, Avishek, et al. Point Pattern Modelling for Degraded Presence-Only Data over Large Regions. Journal of the Royal Statistical Society. Series C (Applied Statistics), vol. 60, no. 5, 2011, pp. 757-776.
 - D. Simpson, J. B. Illian, F. Lindgren, S. H. Sorbye, H. Rue, Going off grid: computationally efficient inference for log-Gaussian Cox processes, Biometrika, 103, 2016, 49-70.



Marked Point Process

▶ A marked point process is composed of a point process and associate marks:

$$\{(s_i, m_i) : i = 1, \ldots, n\}$$

where s_1, \ldots, s_n are locations and m_1, \ldots, m_n are associated marks.

- Example:
 - Simulate from a homogeneous Poisson point process with $\lambda(s)=10$ on $[0,1]^2$
 - Generate $m_i \stackrel{iid}{\sim} Exp(1)$ at each s_i .
 - ▶ The marked point process $\{(s_i, m_i)\}$ is obtained.



A Parametric Model

Ho and Stoyan (2008; Statistics and Probability Letters, 1194-1199):

$$\lambda(s) = exp(\alpha + \beta S(s))$$

$$m(s) = S(s) + \epsilon(s)$$

$$S(s) \sim \mathcal{GP}(\mu(\cdot), C(\cdot, \cdot))$$

That is, using a GP to link m and λ .



- ► There are a lot of interesting research problems related to point process:
 - ▶ Nonstationary proces: estimation of intensity functions
 - Spatio-tempolra point process
 - Relationship between points and marks
- Requiring more work



Summary

► Point process

